

Model based clustering with Missing Not At Random data

Maasai Seminar

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Outline

- 1 Introduction on missing data
- 2 Missing data in model-based clustering
- 3 Estimation of the parameters
- 4 Numerical experiments
- 5 Conclusion

Missing values are everywhere

- Growing masses of data, multiplication of sources
⇒ **Not Available** values (**NA**)
- Our public health application: the **Traumabase[®]** dataset.

250 clinical variables
(heterogeneous)

Trauma.center	Heart rate	Death	Anticoagulant. therapy	Glasgow score	...
Pitie-Salpêtrière	88	0	No	3	
Beaujon	103	0	NA	5	
Bicêtre	NA	0	Yes	6	
Bicêtre	NA	0	No	NA	
Lille	62	0	Yes	6	
Lille	NA	0	No	NA	
⋮	⋮	⋮	⋮	⋮	

1 patient; in total: 30 000 patients

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⋮	⋮	⋮	⋮	⋮	

23 different
hospitals

Missing values are everywhere

Traumabase[®] dataset

- now **30 000** patients.
- **250** heterogeneous variables: continuous, categorical, ordinal,...
- **23** different hospitals
- **missing** values everywhere (1% to 90% NA in each variable).

- **Imputation:** provide a **complete dataset** to the doctors.
- **Estimation:** explain the level of platelet with pre-hospital characteristics.
- **Prediction:** predict the administration or not of the tranexomic acid.
- **Clustering:** identify relevant groups of patients sharing similarities.

Q: *How to deal with missing values?*

What we should not do

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Lille	62	0	Yes	6
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Lille	NA	0	No	NA

Discarding individuals with missing values **is not** a solution

- Loss of information .

Traumabase[®]: only 5% of the rows are kept.

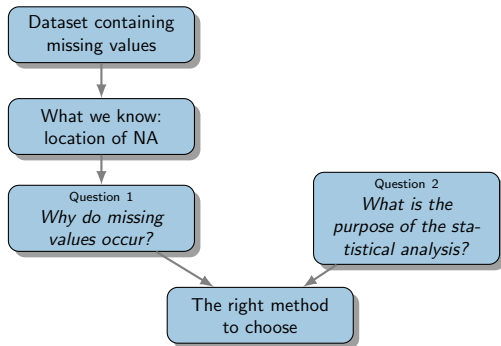
- Bias in the analysis .

Kept observations: sub-population **not necessarily representative** of the overall population.

What we should do: handling missing values

The right method to choose

Q: *How to choose the right method to handle missing values?*



Imputation? Estimation? Prediction?

- The goal is **not necessarily** to obtain a complete dataset.
- A solution can be to **embed missing data management** into the statistical paradigm.

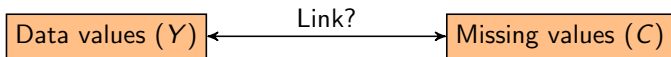
Missing-data notations

- $Y = \{y_1 | \dots | y_n\}^T$: full dataset with n individuals
- Continuous, categorical or mixed data.
- $C = \{c_1 | \dots | c_n\}^T \in \{0, 1\}^{n \times d}$: pattern of missing data for the full dataset

$$c_{ij} = 1 \Leftrightarrow y_{ij} \text{ is missing}$$

- y_i^{obs} : the observed variables values for individual i
- y_i^{mis} : the missing variables values for individual i

Missing-data mechanism (Rubin, 1976)



$$f(c|y; \psi), \psi \in \Omega_\psi$$

Missing Completely At Random (MCAR)

$$f(c|y; \psi) = f(c; \psi)$$

MCAR

Machines fail,
Doctors forget to fill the form

Missing At Random (MAR)

y^{obs} : observed component of y .

$$f(c|y; \psi) = f(c|y^{\text{obs}}; \psi)$$

MAR

Aggregation of datasets

	HR	Death	A. therapy	GCS
Lille	65	0	Yes	6
Lille	59	0	No	4
Pitié	62	0	NA	6
Pitié	84	0	NA	5

Missing Not At Random (MNAR)

The MAR assumption does not hold.
The missingness can depend on the missing data value itself.

MNAR

Emergency situations

HR		HR
65	"underlying" values:	65
59		59
62		62
NA		84

Ignorable vs. non ignorable mechanism

- Parametric estimation: model the joint distribution (Y, C) parametrized by $\gamma, \psi \in \Omega_{\gamma, \psi}$.
- Likelihood-approach: maximizing the full observed likelihood.

$$\begin{aligned}L_{\text{full,obs}}(\gamma, \psi; y^{\text{obs}}, c) &= \int L_{\text{full}}(\gamma, \psi; y, c) dy^{\text{mis}} \\ &= \int f(y; \gamma) f(c|y; \psi) dy^{\text{mis}} \\ &= f(c|y^{\text{obs}}; \psi) \int f(y; \gamma) dy^{\text{mis}} \quad \text{M(C)AR mecha.} \\ &\propto L_{\text{ign}}(\gamma; y^{\text{obs}}) = \int f(y; \gamma) dy^{\text{mis}}\end{aligned}$$

M(C)AR: one can ignore the mechanism.

MNAR: one should consider the mechanism.

Focus on MNAR mechanism

We should consider (Y, C) (not-ignorable mechanism).

The main MNAR specifications

- selection model [Heckman, 1979]:

$$f(y, c; \gamma, \psi) = f(y; \gamma)f(c|y; \psi)$$

- pattern-mixture model [Little, 1993]:

$$f(y, c; \xi, \varphi) = f(c; \xi)f(y|c; \varphi)$$

Q: *How to choose the MNAR specification ?*

- Estimate the parameters of the data distribution: **selection models**.
- Distribution is not the same for the observed data and the missing data: pattern-mixture models.

Focus on MNAR mechanism

We should prove the identifiability of the parameters.

Identifiability issue in the MNAR case Credit: Ilya Shpitser

$$Y^{\text{NA}} = [1, \text{NA}, 0, 1, \text{NA}, 0].$$

- **Case 1:** Y missing only if $Y = 1$.

$$Y = [1, 1, 0, 1, 1, 0], \mathbb{P}(Y = 1) = 2/3.$$

- **Case 2:** Y missing only if $Y = 0$.

$$Y = [1, 0, 0, 1, 0, 0], \mathbb{P}(Y = 1) = 1/3.$$

\Rightarrow We start from 2 equal observed distribution. It leads to different parameters of the data distribution $\mathbb{P}(Y = 1)$.

Identifiability: the parameters of (Y, C) are uniquely determined from available information $(Y, C = 0)$.

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Our goals

- MNAR mechanism.
- Selection model: $f(y, c; \gamma, \psi) = f(y; \gamma)f(c|y; \psi)$.

Embed missing data management into the analysis to:

- Perform clustering: identify relevant groups of individuals.
- Estimate the parameters of the data distribution.
- (Impute missing values.)

Clustering: model-based approach

- Partition with K clusters: $Z = (z_1 | \dots | z_n)^T \in \{0, 1\}^{n \times K}$, with $z_{ik} = 1$ if y_i belongs to cluster k .

Mixture model

$$f(y_i; \gamma) = \sum_{k=1}^K \underbrace{\mathbb{P}(z_{ik}=1)}_{\pi_k} \underbrace{f_k(y_i; \lambda_k)}_{\text{pdf in the cluster } k}$$

Clustering: model-based approach

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Mixture model

$$f(y_i; \gamma) = \sum_{k=1}^K \underbrace{\pi_k}_{=\mathbb{P}(z_{ik}=1)} \underbrace{f_k(y_i; \lambda_k)}_{\text{pdf in the cluster } k}$$

- Missing data in Y .

Mixture model with missing data

$$f(y_i, c_i; \theta) = \sum_{k=1}^K \pi_k f_k(y_i; \lambda_k) f_k(c_i | y_i; \psi_k),$$

A zoology of MNAR models in clustering

Q: Which distribution $f_k(c_i | y_i; \psi_k)$ to propose in this clustering context?

$$f_k(c_i | y_i; \psi_k) = \prod_{j=1}^d (\rho(\alpha_{kj} + \beta_{kj}y_{ij}))^{c_{ij}} (1 - \rho(\alpha_{kj} + \beta_{kj}y_{ij}))^{1-c_{ij}},$$

where $\psi_k = (\alpha_{k1}, \beta_{k1}, \dots, \alpha_{kK}, \beta_{kK})$ and ρ is a link function.

How to understand this distribution?

- α_{kj} : the missingness depends on the class membership k , not the same effect for every variable.
- β_{kj} : the missingness depends on the value itself (y_{ij}), not the same effect for each cluster.
- Simplest model:

$$\text{MCAR: } \beta_{kj} = 0, \forall (k, j) \text{ and } \alpha_{1j} = \dots = \alpha_{Kj}, \forall j.$$

A zoology of MNAR models in clustering

Parcimonious models: the probability of being missing depend

- **on both the variable and the class membership:**

$$\text{MNAR}_{yz^j}: \beta_{1j} = \dots = \beta_{Kj}, \forall j.$$

$$\text{MNAR}_{y^kz}: \alpha_{kj} = \dots = \alpha_{k1}, \forall k.$$

$$\text{MNAR}_{yz}: \beta_{1j} = \dots = \beta_{Kj}, \forall j \text{ and } \alpha_{kj} = \dots = \alpha_{k1}, \forall k.$$

- **only on the variable itself:**

$$\text{MNAR}_y: \alpha_{11} = \dots = \alpha_{1d} = \alpha_{21} = \dots = \alpha_{Kd} \text{ and } \beta_{1j} = \dots = \beta_{Kj}, \forall j.$$

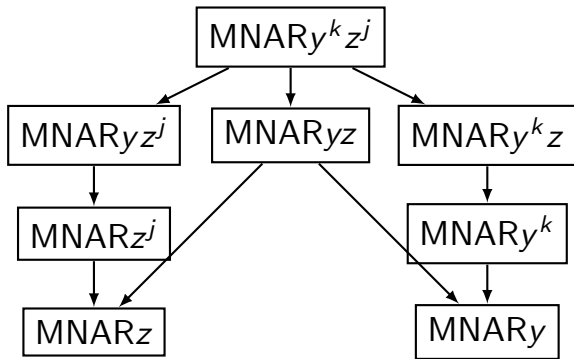
$$\text{MNAR}_{y^k}: \alpha_{11} = \dots = \alpha_{1d} = \alpha_{21} = \dots = \alpha_{Kd}.$$

- **only on the class membership:**

$$\text{MNAR}_z: \beta_{kj} = 0, \forall (k, j) \text{ and } \alpha_{kj} = \dots = \alpha_{kd}, \forall k.$$

$$\text{MNAR}_{z^j}: \beta_{kj} = 0, \forall (k, j).$$

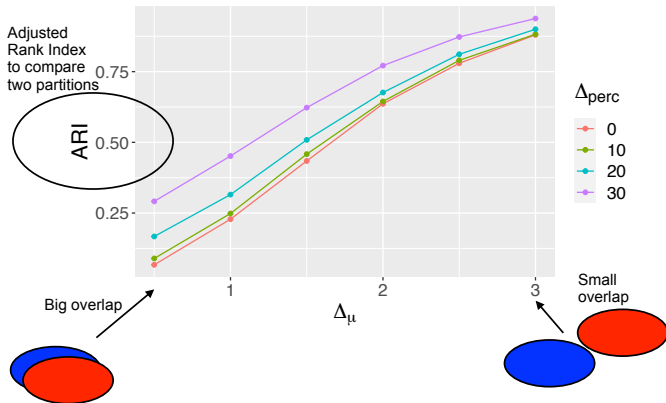
Proposed zoology of MNAR models in clustering



MNAR_Z from every angle

(1) C gives information on partition Z

- MNAR_Z model, Bivariate Gaussian model
- cluster overlap: $\Delta_{\mu} = |\mu_1 - \mu_2|$ varies.
- difference of percentage of NA between the 2 clusters: Δ_{perc} varies.



MNAR_Z from every angle

(2) MNAR_Z (and MNAR_Z^j) models interpreted as MAR

$$Y^{\text{obs}} = \begin{pmatrix} ? & 2.6 & 5 \\ \text{blue} & 1.9 & 4 \\ \text{red} & 2.3 & ? \end{pmatrix}, \quad C = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

$$\tilde{Y}^{\text{obs}} = \begin{pmatrix} ? & 2.6 & 5 & 1 & 0 & 0 \\ \text{blue} & 1.9 & 4 & 0 & 0 & 0 \\ \text{red} & 2.3 & ? & 0 & 0 & 1 \end{pmatrix}.$$

Proposition 1: in terms of maximum likelihood

MLE associated to \tilde{Y}^{obs} under MAR model
 \Leftrightarrow MLE associated to Y^{obs} under MNAR_Z/MNAR_Z^j models.

Identifiability results

Previous works: [Teicher, 1963], [Allman et al., 2009] (without NA), [Miao et al., 2016] (for MNAR data).

Proposition 2: identifiability for continuous and count data

Assume

- 1 The marginal mixture $\sum_{k=1}^K \pi_k f_k(y_i; \theta_k)$ is identifiable
- 2 There exists a total ordering \preceq of $\mathcal{F}_j \times \mathcal{R}$, for $j \in \{1, \dots, d\}$ fixed, where $\mathcal{F}_j = \{f_{1j}, \dots, f_{Kj}\}$ and $\mathcal{R} = \{\rho_1, \dots, \rho_K\}$.

The mixture model with any MNAR* is identifiable.

Proposition 3: identifiability for categorical data

Assume $d_{\text{cat}} \geq 2 \lceil \log_2 K \rceil + 1$ and $f_k(\cdot; \theta_k) = \prod_{j=1}^d f_{kj}(\cdot; \theta_{kj})$

- ✓ The mixture model with MNARz or MNARz^j is identifiable.
- ✗ The mixture model with any MNARy* is not identifiable.

- For mixed data: result follows from Proposition 2 and 3.

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EM algorithm

Initialized at the point $\theta^{[0]} = (\pi^{[0]}, \lambda^{[0]}, \psi^{[0]})$, the iteration $[r]$ of the EM algorithm consists in performing two steps:

- **E-step**: compute the expectation of the complete-data log-likelihood $Q(\theta; \theta^{[r-1]}) = \mathbb{E}_{\theta^{[r-1]}} [\ell_{\text{comp}}(\theta; Y, Z, C) \mid Y^{\text{obs}}, C]$,
$$\ell_{\text{comp}}(\theta; Y, Z, C) = \sum_{i=1}^n \log \left(\sum_{k=1}^K \pi_k f_k(y_i; \lambda_k) f_k(c_i \mid y_i; \psi_k) \right).$$
- **M-step**: update the parameters by maximizing this function
$$\theta^{[r]} = \operatorname{argmax}_{\theta} Q(\theta; \theta^{[r-1]}).$$

EM algorithm: feasible computations?

One has: $Q(\theta; \theta^{[r-1]}) =$

$$\sum_i \sum_k t_{ik}(\theta^{[r-1]}) \left[\log(\pi_k) + \underbrace{\tau_y(\lambda_k; y_i^{\text{obs}}, c_i, \theta^{[r-1]})}_{=\mathbb{E}_{\theta^{[r-1]}}[\ln f_k(y_i; \lambda_k) | y_i^{\text{obs}}, c_i, z_{ik}=1]} + \underbrace{\tau_c(\psi_k; y_i^{\text{obs}}, c_i, \theta^{[r-1]})}_{\mathbb{E}_{\theta^{[r-1]}}[\ln f_k(c_i | y_i; \phi_k) | y_i^{\text{obs}}, c_i, z_{ik}=1]} \right]$$

with $t_{ik}(\theta^{[r-1]}) = \mathbb{P}(z_{ik} = 1 | y_i^{\text{obs}}, c_i)$.

- Law of y_i^{mis} given $(y_i^{\text{obs}}, z_{ik} = 1, c_i)$?
- Computation of the expectation over this law of $f_k(c_i | y_i; \phi_k)$?

EM algorithm: feasible computations?

MNAR_z, MNAR_{zj}: needs some computations but still simple.

$$f_k(c_i | y_i; \psi_k) = \rho(\alpha_{kj}) \quad (\perp\!\!\!\perp Y)$$

- $\Rightarrow \mathcal{L}(y_i^{\text{mis}} | y_i^{\text{obs}}, z_{ik} = 1, c_i) = \mathcal{L}(y_i^{\text{mis}} | y_i^{\text{obs}}, z_{ik} = 1)$
- EM algorithm for Gaussian data,
- EM for categorical data.

MNAR_{y*}: needs approximations

$$f_k(c_i | y_i; \psi_k) = \rho(\alpha_{kj} + \beta_{kj} y_{ij}) \quad (\text{not } \perp\!\!\!\perp Y)$$

- $(y_i^{\text{mis}} | y_i^{\text{obs}}, z_{ik} = 1, c_i)$ not classical if Logit link.
- No closed forms.

SEM algorithm for MNAR_y*

SEM easier? random drawing instead of expectation

- **SE-step**: draw the missing data

$$((y_i^{\text{mis}})^{[r]}, z_i^{[r]}) \sim (\cdot \mid y_i^{\text{obs}}, c_i; \pi^{[r-1]}, \lambda^{[r-1]}, \psi^{[r-1]})$$

- $(y_i^{\text{mis}})^{[r]} \sim (\cdot \mid y_i^{\text{obs}}, z_i^{[r-1]}, c_i; \lambda^{[r-1]}, \psi^{[r-1]})$:

- $z_i^{[r]} \sim (\cdot \mid y_i^{[r]}, c_i; \pi^{[r-1]}, \lambda^{[r-1]}, \psi^{[r-1]})$: draw the membership k of $z_i^{[r]}$ from the **multinomial distribution**

Let $Y^{[r]} = (y_1^{[r]} \mid \dots \mid y_n^{[r]})$, $Z^{[r]} = (z_1^{[r]} \mid \dots \mid z_n^{[r]})$ be the imputed matrix and the partition.

- **M-step**: for $k = 1, \dots, K$, compute $\pi_k^{[r]}, \lambda_k^{[r]}, \psi_k^{[r]}$.

$$(\cdot \mid y_i^{\text{obs}}, z_i^{[r-1]}, c_i; \lambda^{[r-1]}, \psi^{[r-1]})?$$

- not classical if ρ is Logit
- truncated Gaussian distribution if ρ is Probit

Summary of the algorithms

	EM		SEM	
	Gaussian	Categorical	Gaussian	Categorical
MNAR _z MNAR _{z^j}	✓	✓	✓	✓
MNAR _{y*}	no closed form	not ident.	✓ (Probit)	not ident.

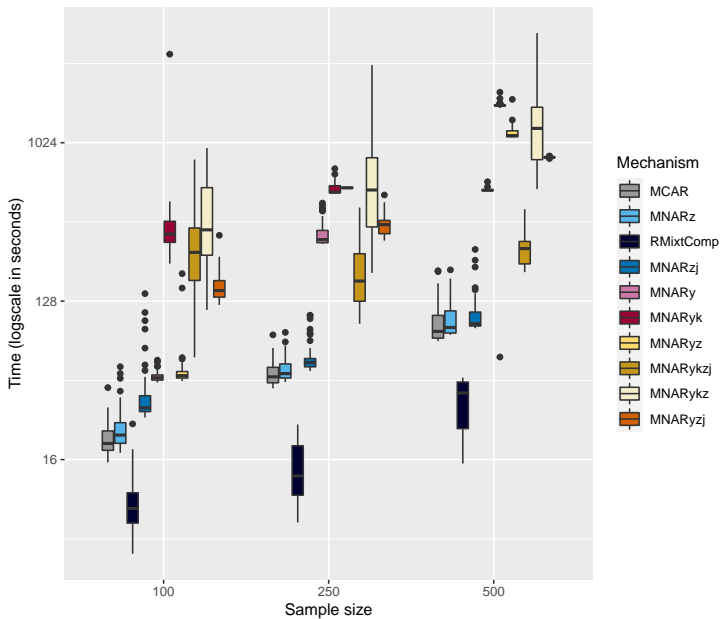
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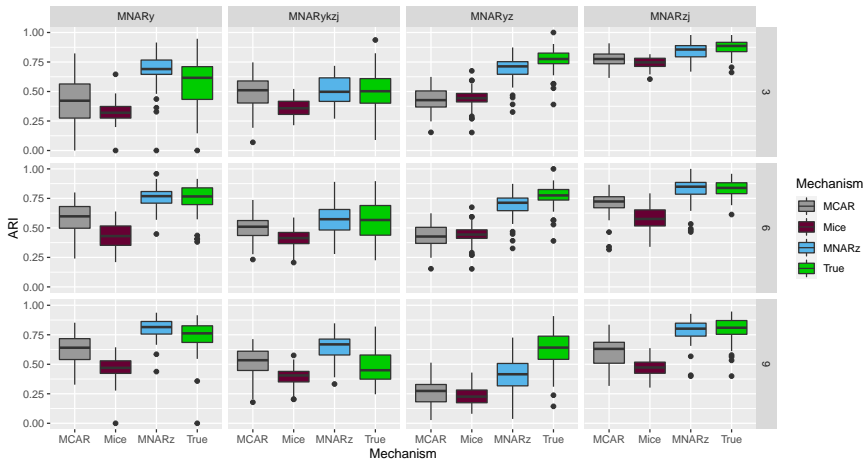
Setting

- Gaussian mixture with three components having unequal proportions ($1 = 0.5, 2 = 3 = 0.25$), independent variables.
- Control the rates of misclassification (10%) and missingness (30%): we fix them equal for each scenario.

Computation time

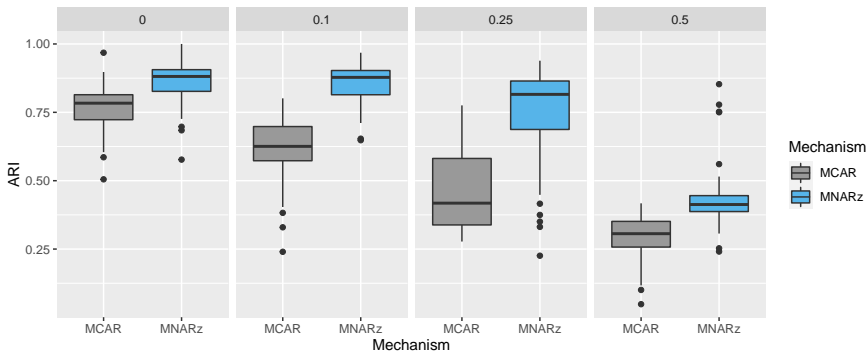


MNAR_Z vs other MNAR model



MNAR_z: robustness to the misspecification of the data distribution

- Three-components Gaussian mixture with non-diagonal covariance matrices: $\Sigma_{ij} = \ell, i \neq j$, with $\ell \in \{0, 0.1, 0.25, 0.5\}$
- Algorithm assumes $\ell = 0$.

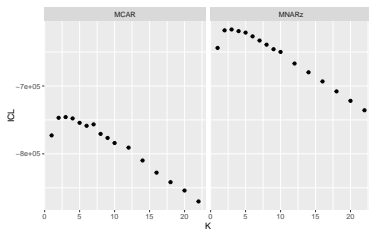


Results on real data

41 mixed variables containing missing values assumed to be MNARz. The variables related to the patient death are not taken into account.

Can the MNAR mechanism improve the classification ? Is there an influence of the mechanism ?

- Same number of clusters selected by the ICL criterion;
- For $K = 3$, ARI between the classifications obtained assuming MNARz and MCAR = 0.9;



Results on rela data

$$\sqrt{\sum_{i=1}^n (\mathbb{P}(z_{ik} = 1 | y_{is}^{\text{obs}}; \theta^{\text{MCAR}}) - \mathbb{P}(z_{i\tilde{k}} = 1 | y_{is}^{\text{obs}}; \theta^{\text{MNARz}}))^2, \forall k, \tilde{k} \in \{1, 2, 3\}}$$

	MNARz	Class 1	Class 2	Class 3
MCAR				
Class 1		2.43	26.5	37.6
Class 2		26.2	3.40	20.1
Class 3		39.3	19.2	2.05

Table: Euclidean distance between the conditional probabilities of the cluster memberships given the observed values of the variable Shock.index.ph in the Traumabase dataset, obtained using the algorithm considering MNARz data, and the ones obtained with the algorithm considering MCAR data.

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Conclusion

Summary

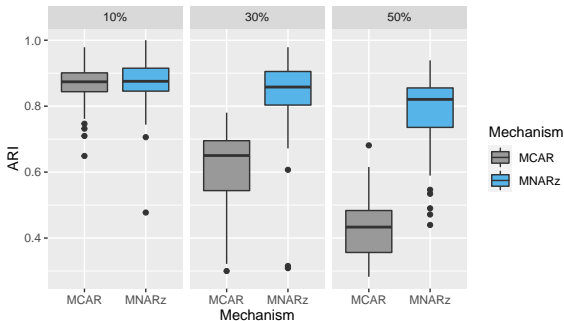
- Interest to put a model on c
- Interest of the simple but meaningful model MNAR_z
- Trade-off between biased mixture model and biased missingness mechanism.

Ongoing works

- Implement the proposed models/algo. in the Mixmod software^a





^a<http://www.mixmod.org>

MNAR_z: robustness to the NA% and choice of K



	MCAR	MNARz
10% NA	94%	94%
30% NA	8%	56%
50% NA	0%	20%

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