# Estimation and imputation in Probabilistic Principal Component Analysis with Missing Not At Random data





### **Proposal:**

Handling several MNAR variables (coupled with M(C)AR variables) in PPCA model without modeling the missing-data mechanism and using only the observed information: identifi**ability** and **estimation** of the model parameters and **imputation** of the missing values.

### Missing data

- One of the ironies of working with Big Data is that missin data plays an ever more significant role.
- Three types of **missing-data mechanisms** [1]:
- MCAR missing values does not depend on the data.
- MAR missing values depends on the observed variables.
- MNAR missing values depends on both observed and unobserv data such as its value itself.
- Most methods focus on the easiest M(C)AR data, here focus on MNAR data.

# Existing works for MNAR data

- Modeling the MNAR mechanism [2, 3]. × Parametric assumption for the mechanism distribution × Computationally costly.
- Without modeling the mechanism and by only using a available observed cells [4, 5, 6].
- × Restricted to simple linear models with few missing var
- Most of the works consider **self-masked MNAR** var ables: the missingness of a variable depends on the var able itself. E.g. the probability to have a missing value o income depends on the value of income (rich people les inclined to reveal their income).

### References

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Setting	5)
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Setting	
<ul> <li>Data matrix Y ∈ ℝ<sup>n×p</sup>,</li> <li>Coefficients matrix B ∈ ℝ<sup>r×p</sup> of rank r &lt; min{n, p}</li> <li>r latent variables grouped in W ∈ ℝ<sup>n×r</sup>,</li> <li>Ω ∈ ℝ<sup>n×p</sup> the missing-data pattern: Ω<sub>ij</sub> = 1 if Y<sub>ij</sub> is observed, 0 otherwise.</li> </ul>	$ \begin{array}{c} & \dashrightarrow & A \\ Y \odot \Omega \\ & Y \odot \Omega \\ & & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ &$
PPCA model	
$Y = 1\alpha + WB + \epsilon, \text{ with } \begin{cases} W = (W_{1.}  \dots  W_n] \\ \alpha \in \mathbb{R}^p \text{ and } 1 = (1) \\ \epsilon = (\epsilon_{1.}  \dots  \epsilon_{n.})^T, \end{cases}$	$(.)^T, W_i$ $\sim$ $(1)^T \in$ $\epsilon_{i.} \sim \mathcal{N}(0)$
PPCA model identifiability w	ith M
<ul> <li>Identifying the PPCA model</li> <li>⇔ identifying the missing-data mechanism.</li> <li>Assumptions:</li> </ul>	
<b>A01.</b> $d$ self-masked MNAR var. and $p - d$ of var., $F_m$ known strictly monotone with $\mathbb{P}(\Omega_{im} = 1   Y_{i.}) = F_m(\phi_m^0 + \phi_m^1 Y_{im}), T_m$ missing-data mechanism.	ther MCA a finite s with $\phi_m$
A02. $\forall (k, \ell) \in \{1, \dots, p\}^2,  k \neq \ell, \qquad \Omega_{.k}$	$\square \Omega_{\ell}   Y$
Proposition 1: ident	lability
<ul> <li>Under A01. and A02., the PPCA parmissing-data mechanism parameter φ a</li> <li>Assuming that the noise level σ<sup>2</sup> is known trix B is identifiable up to a row permutation.</li> </ul>	ameters re identi wn, the o tation.
General MNAR sett estimation/imputa	ting for ation

• r pivot variables indexed by  $\mathcal{J}$  observed or MCAR. • d general MNAR variables indexed by  $\mathcal{M}$ , missingness depends on all the variables except r ones. With  $\mathcal{J} = \{1, \ldots, p\} \setminus \mathcal{J}$ ,  $\forall m \in \mathcal{M}, \quad \mathbb{P}(\Omega_{im} = 1 | Y_{i.}) = \mathbb{P}(\Omega_{im} = 1 | (Y_{ik})_{k \in \overline{\mathcal{I}}}).$ 

## Estimation with MNAR data





iNAR data

AR (or observed) support  $= (\phi_m^1, \phi_m^2)$  the

# $(\alpha, \Sigma)$ and the fiable. coefficient ma-

- Toy exemple: p = 3, r = 2,  $Y_{.1}$  MNAR,  $Y_{.2}$ ,  $Y_{.3}$  observed.
- $(Y_{.1} Y_{.2} Y_{.3}) = \mathbf{1} (\alpha_1 \alpha_2 \alpha_3) + (W_{.1} W_{.2}) B + \epsilon.$
- Assumption: fully connected PPCA i.e. any variable generated by all the latent variables  $\Rightarrow$  linear links can be established.

### Assumptions:

- A1.  $(B_{.1} B_{.2})$  is invertible
- A2.  $Y_{.2} \perp \Omega_{.1} | Y_{.3}$
- **A3.** Consistent estimators for  $\alpha_2$  and  $\alpha_3$
- A4. Consistent estimators for  $(\mathcal{B}_{2\to 1,3[k]}^c)_{k\in\{0,1,3\}}$

### **Proposition 4:** mean estimation

- Under A1. and A2., one defines the estimator for the mean of  $Y_{:1}$ • Under A3. and A4.,  $\hat{\alpha}_1$  is a consistent estimator of  $\alpha_1$ .
- Same method for the variance and covariances,  $\hat{\Sigma}$  estimates  $\Sigma$ .
- $\hat{\Sigma} \sigma^2 \mathrm{Id}_{3\times 3}$  estimates  $B^T B \Rightarrow$  estimation of B  $(r, \sigma^2 \text{ known})$ .

### Imputation of the missing values

tional expectation of  $(Y_{i1})$  given  $Y_{i2}$  and  $Y_{i3}$ .

- Pivot variable selection: with experts or var. with less %NA, bigger set (> r) and estimates aggregation.
- and  $Y_{:3}$ . •  $(\mathcal{B}_{2\to 1,3[k]}^c)_{k\in\{0,1,3\}}$  estimated by the coefficients of the linear regression of  $Y_{.2}$  on  $Y_{.1}$  and  $Y_{.3}$  using the rows where  $Y_{.1}$  is • CV strategy to estimate  $\sigma^2$  and r. observed.

# Application to clinical data TraumaBase<sup>®</sup>

- **EMMAR**: EM algorithm to perform PPCA with MAR values [7].
- SoftMAR: matrix completion using iterative SVD algorithm for M(C)AR values [8].
- **MNARparam**: low-rank method for MNAR values (modeling the mechanism) [3].
- Mean: naive imputation by the mean.





 $\checkmark$  fully connected PPCA  $\checkmark$  MNAR mechanism  $\checkmark$  observed variables  $\checkmark$  noise tends to zero

 $\hat{\alpha}_{1} := \frac{\hat{\alpha}_{2} - \hat{\mathcal{B}}_{2 \to 1,3[0]}^{c} - \hat{\mathcal{B}}_{2 \to 1,3[3]}^{c} \hat{\alpha}_{3}}{\hat{\mathcal{B}}_{2 \to 1,3[1]}^{c}}, \ (\mathcal{B}_{2 \to 1,3[k]}^{c})_{k \in \{0,1,3\}}: \ effects \ of \ Y_{.2} \ on \ Y_{.1}, Y_{.3} \ when \ \Omega_{.1} = 1.$ 

• Impute the missing values  $Y_{i1}$  for  $i \in \{1, \ldots, n\}$  such that  $M_{i1} = 0$  using the condi-

### **Practical implementation**

•  $\hat{\alpha}_2$  and  $\hat{\alpha}_3$  are computed as empirical means of  $Y_{:2}$ 



Figure 1: Our method **MNAR** compared with others methods in terms of imputation error.